

MODULE 7**LIMITING DISTRIBUTIONS****LECTURE 41****Topics****7.3 SOME PRESERVATION RESULTS****7.3 SOME PRESERVATION RESULTS**

In this section, we will discuss the algebraic operations under which convergence in probability and/or convergence in distribution is preserved.

Theorem 3.1

Let $\{X_n\}_{n \geq 1}$ and $\{Y_n\}_{n \geq 1}$ be sequences of random variables and let X be another random variable.

- (i) Let $g: \mathbb{R} \rightarrow \mathbb{R}$ be continuous at $c \in \mathbb{R}$ and let $X_n \xrightarrow{p} c$, as $n \rightarrow \infty$. Then $g(X_n) \xrightarrow{p} g(c)$, as $n \rightarrow \infty$.
- (ii) Let $h: \mathbb{R}^2 \rightarrow \mathbb{R}$ be continuous at $(c_1, c_2) \in \mathbb{R}^2$ and let $X_n \xrightarrow{p} c_1, Y_n \xrightarrow{p} c_2$, as $n \rightarrow \infty$. Then $h(X_n, Y_n) \xrightarrow{p} h(c_1, c_2)$, as $n \rightarrow \infty$.
- (iii) Let $g: \mathbb{R} \rightarrow \mathbb{R}$ be continuous on a support S_X of X and let $X_n \xrightarrow{d} X$, as $n \rightarrow \infty$. Then $g(X_n) \xrightarrow{d} g(X)$, as $n \rightarrow \infty$.
- (iv) Let $h: \mathbb{R}^2 \rightarrow \mathbb{R}$ be continuous at all points in $D = \{(x, b): x \in S_X\}$, where b is a fixed real constant and S_X is a support of X . If $X_n \xrightarrow{d} X$ and $Y_n \xrightarrow{p} b$, as $n \rightarrow \infty$, then $h(X_n, Y_n) \xrightarrow{d} h(X, b)$, as $n \rightarrow \infty$.

Proof. We shall not attempt to prove assertions (iii) and (iv) here as their proofs are slightly involved.

- (i) Fix $\varepsilon > 0$. Since $g: \mathbb{R} \rightarrow \mathbb{R}$ is continuous at $c \in \mathbb{R}$, there exists a $\delta \equiv \delta(\varepsilon, c)$ such that

$$|x - c| < \delta \Rightarrow |g(x) - g(c)| < \varepsilon.$$

or equivalently

$$|g(x) - g(c)| \geq \varepsilon \Rightarrow |x - c| \geq \delta.$$

Therefore,

$$\begin{aligned} 0 &\leq P(\{|g(X_n) - g(c)| \geq \varepsilon\}) \leq P(\{|X_n - c| \geq \delta\}) \xrightarrow{n \rightarrow \infty} 0 \quad (\text{since } X_n \xrightarrow{p} c) \\ \Rightarrow \lim_{n \rightarrow \infty} P(\{|g(X_n) - g(c)| \geq \varepsilon\}) &= 0 \\ \Rightarrow g(X_n) &\xrightarrow{p} g(c), \quad \text{as } n \rightarrow \infty. \end{aligned}$$

- (ii) Fix $\varepsilon > 0$. Since $h: \mathbb{R}^2 \rightarrow \mathbb{R}$ is continuous at $(c_1, c_2) \in \mathbb{R}^2$, there exists a $\delta = \delta(\varepsilon, c_1, c_2)$ such that

$$|x - c_1| < \delta \text{ and } |y - c_2| < \delta \Rightarrow |h(x, y) - h(c_1, c_2)| < \varepsilon,$$

or equivalently

$$|h(x, y) - h(c_1, c_2)| \geq \varepsilon \Rightarrow |x - c_1| \geq \delta \text{ or } |y - c_2| \geq \delta.$$

Therefore,

$$\begin{aligned} P(\{|h(X_n, Y_n) - h(c_1, c_2)| \geq \varepsilon\}) &\leq P(\{|X_n - c_1| \geq \delta\} \cup \{|Y_n - c_2| \geq \delta\}) \\ &\leq P(\{|X_n - c_1| \geq \delta\}) + P(\{|Y_n - c_2| \geq \delta\}) \quad (\text{using Boole's inequality}) \\ &\xrightarrow{n \rightarrow \infty} 0 + 0 = 0 \quad (\text{since } X_n \xrightarrow{p} c_1 \text{ and } Y_n \xrightarrow{p} c_2) \\ \Rightarrow \lim_{n \rightarrow \infty} P(\{|h(X_n, Y_n) - h(c_1, c_2)| \geq \varepsilon\}) &= 0 \\ \Rightarrow h(X_n, Y_n) &\xrightarrow{p} h(c_1, c_2), \text{ as } n \rightarrow \infty. \blacksquare \end{aligned}$$

Throughout, we shall use the following convention. If, for a real constant c , we write $X_n \xrightarrow{d} c$, as $n \rightarrow \infty$, then it would mean that X_n converges in distribution, as $n \rightarrow \infty$, to a random variable degenerate at c (i.e., $X_n \xrightarrow{p} c$, as $n \rightarrow \infty$). Similarly, for a random variable X , $0 \times X$ will be treated as a random variable degenerate at 0.

Now we provide the following useful lemma whose proof, being straight forward, is left as an exercise.

Lemma 3.1

- (i) Let X and Y be random variables and let c be a real constant. If $P(\{Y = c\}) = 1$ then $X + Y \stackrel{d}{=} X + c$ and $XY \stackrel{d}{=} cX$, where $0 \times X$ is treated as a random variable degenerate at 0.
- (ii) Let $\{X_n\}_{n \geq 1}$ and $\{Y_n\}_{n \geq 1}$ be sequences of real numbers such that $X_n \stackrel{d}{=} Y_n, n = 1, 2, \dots$. If, for some real constant $c, X_n \xrightarrow{p} c$, as $n \rightarrow \infty$, then $Y_n \xrightarrow{p} c$, as $n \rightarrow \infty$.
- (iii) Let $\{X_n\}_{n \geq 1}$ and $\{Y_n\}_{n \geq 1}$ be sequences of real numbers such that $X_n \stackrel{d}{=} Y_n, n = 1, 2, \dots$. If, for some random variable $X, X_n \xrightarrow{d} X$, as $n \rightarrow \infty$, then $Y_n \xrightarrow{d} X$, as $n \rightarrow \infty$.
- (iv) Let $\{a_n\}_{n \geq 1}$ be sequence of real numbers such that $\lim_{n \rightarrow \infty} a_n = a \in \mathbb{R}$ and let $\{X_n\}_{n \geq 1}$ be a sequence of random variables such that X_n is degenerate at $a_n, n = 1, 2, \dots$. Then $X_n \xrightarrow{p} a$, as $n \rightarrow \infty$. ■

Theorem 3.2

Let $\{X_n\}_{n \geq 1}$ and $\{Y_n\}_{n \geq 1}$ be sequences of random variables and let $\{a_n\}_{n \geq 1}$ and $\{b_n\}_{n \geq 1}$ be sequences of real numbers such that $\lim_{n \rightarrow \infty} a_n = a$ and $\lim_{n \rightarrow \infty} b_n = b$.

- (i) Suppose that, for some real constants c_1 and $c_2, X_n \xrightarrow{p} c_1$ and $Y_n \xrightarrow{p} c_2$, as $n \rightarrow \infty$. Then, as $n \rightarrow \infty, X_n + Y_n \xrightarrow{p} c_1 + c_2, X_n - Y_n \xrightarrow{p} c_1 - c_2$ and $X_n Y_n \xrightarrow{p} c_1 c_2$. Moreover, if $c_2 \neq 0$, then $\frac{X_n}{Y_n} \xrightarrow{p} \frac{c_1}{c_2}$, as $n \rightarrow \infty$.
- (ii) Suppose that, for a real constant c and a random variable $X, X_n \xrightarrow{d} X$ and $Y_n \xrightarrow{p} c$, as $n \rightarrow \infty$. Then, as $n \rightarrow \infty, X_n + Y_n \xrightarrow{d} X + c, X_n - Y_n \xrightarrow{d} X - c$ and $X_n Y_n \xrightarrow{d} cX$. Moreover, if $c \neq 0$, then $\frac{X_n}{Y_n} \xrightarrow{d} \frac{X}{c}$, as $n \rightarrow \infty$.
- (iii) Suppose that, for a real constant $c, X_n \xrightarrow{p} c$, as $n \rightarrow \infty$. Then $a_n X_n + b_n \xrightarrow{p} ac + b$, as $n \rightarrow \infty$.
- (iv) Suppose that, for a random variable $X, X_n \xrightarrow{d} X$, as $n \rightarrow \infty$. Then $a_n X_n + b_n \xrightarrow{d} aX + b$, as $n \rightarrow \infty$.

Proof. (i) and (ii) follow from Theorem 3.1 (ii) and (iv) as $h_1(x, y) = x + y, h_2(x, y) = x - y$ and $h_3(x, y) = xy$ are continuous functions on \mathbb{R}^2 , and $h_4(x, y) = \frac{x}{y}$ is continuous on $D = \{(s, t) \in \mathbb{R}^2 : t \neq 0\}$.

- (iii) Let Y_n be a random variable that is degenerate at a_n and let Z_n be a random variable that is degenerate at $b_n, n = 1, 2, \dots$. Then $Y_n \xrightarrow{p} a$ and $Z_n \xrightarrow{p} b$, as $n \rightarrow \infty$

- (Lemma 3.1 (iv)). Now using (i) we get $X_n Y_n + Z_n \xrightarrow{p} ac + b$, as $n \rightarrow \infty$. Since $a_n X_n + b_n \stackrel{d}{=} X_n Y_n + Z_n$, $n = 1, 2, \dots$, (Lemma 3.1 (i)), the assertion follows on using Lemma 3.1 (ii).
- (iv) Let Y_n and Z_n be as defined in (iii). Then $Y_n \xrightarrow{p} a$ and $Z_n \xrightarrow{p} b$, as $n \rightarrow \infty$. Using (ii) we get $X_n Y_n + Z_n \xrightarrow{d} aX + b$, as $n \rightarrow \infty$. Since $a_n X_n + b_n \stackrel{d}{=} X_n Y_n + Z_n$, $n = 1, 2, \dots$, the assertion follows on using Lemma 3.1(iii). ■

Remark 3.1

The CLT asserts that if X_1, X_2, \dots are i.i.d. random variables with mean μ and finite variance $\sigma^2 > 0$, then

$$Z_n \stackrel{\text{def}}{=} \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{d} Z \sim N(0,1), \text{ as } n \rightarrow \infty,$$

where $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$. Since $\frac{\sigma}{\sqrt{n}} \rightarrow 0$, as $n \rightarrow \infty$, using Theorem 3.2 (iv) we get

$$\bar{X}_n - \mu = \frac{\sigma}{\sqrt{n}} Z_n \xrightarrow{d} 0 \times Z, \text{ as } n \rightarrow \infty.$$

Note that $0 \times Z$ is a random variable degenerate at 0. Thus it follows that

$$\begin{aligned} \bar{X}_n - \mu &\xrightarrow{d} 0, & \text{as } n \rightarrow \infty \\ \Leftrightarrow \bar{X}_n - \mu &\xrightarrow{p} 0, & \text{as } n \rightarrow \infty \\ \Leftrightarrow \bar{X}_n &\xrightarrow{p} \mu, & \text{as } n \rightarrow \infty. \end{aligned}$$

The above discussion suggests that, under the finiteness of second moment (or variance), the CLT is a stronger result than the WLLN. ■

Example 3.1

Let $\{X_n\}_{n \geq 1}$ and $\{Y_n\}_{n \geq 1}$ be sequences of random variables.

- (i) If $X_n \xrightarrow{p} \ln 4$ and $Y_n \xrightarrow{p} 2$, as $n \rightarrow \infty$, show that $X_n + \ln Y_n \xrightarrow{p} \ln 8$ and $e^{X_n} \ln Y_n \xrightarrow{p} \ln 16$, as $n \rightarrow \infty$;
- (ii) If $X_n \xrightarrow{d} Z \sim N(0,1)$, as $n \rightarrow \infty$, show that $X_n^2 \xrightarrow{d} Q_1 \sim \chi_1^2$ (the chi-square distribution with one degree of freedom), as $n \rightarrow \infty$.
- (iii) If $X_n \xrightarrow{d} Z \sim N(0,1)$, and $Y_n \xrightarrow{p} 3$, as $n \rightarrow \infty$, show that $X_n Y_n \xrightarrow{d} V \sim N(0, 9)$ and $2X_n + 3Y_n \xrightarrow{d} Q_2 \sim N(9, 4)$, as $n \rightarrow \infty$.

- (iv) For a given $\theta > 0$, if X_1, X_2, \dots are i.i.d. $U(0, \theta)$ random variables and $X_{n:n} = \max\{X_1, \dots, X_n\}$, $n = 1, 2, \dots$, show that $e^{X_{n:n}} \xrightarrow{p} e^\theta$, $X_{n:n}^2 + X_{n:n} + 1 \xrightarrow{p} \theta^2 + \theta + 1$ and $e^{-\frac{n(\theta - X_{n:n})}{\theta}} \xrightarrow{d} U \sim U(0, 1)$, as $n \rightarrow \infty$.

Solution.

- (i) Since $h_1(x) = \ln x$, $x \in (0, \infty)$ is a continuous function, using Theorem 3.1 (i) it follows that $\ln Y_n \xrightarrow{p} \ln 2$, as $n \rightarrow \infty$. Now on using Theorem 3.2 (i) we get $X_n + \ln Y_n \xrightarrow{p} \ln 4 + \ln 2 = \ln 8$, as $n \rightarrow \infty$. Also, since $h_2(x) = e^x$, $x \in \mathbb{R}$, is a continuous function on \mathbb{R} , on using Theorem 3.1 (i), we get $e^{X_n} \xrightarrow{p} e^{\ln 4} = 4$, as $n \rightarrow \infty$. Now on using Theorem 3.2(i) it follows that $e^{X_n} \ln Y_n \xrightarrow{p} 4 \ln 2 = \ln 16$, as $n \rightarrow \infty$.
- (ii) Since $h_3(x) = x^2$, $x \in \mathbb{R}$, is a continuous function on \mathbb{R} , using Theorem 3.1 (iii) we get $X_n^2 \xrightarrow{d} Z^2$, as $n \rightarrow \infty$. Let $Q_1 = Z^2$. Since $Z \sim N(0, 1)$, we have $Q_1 \sim \chi_1^2$ (Theorem 4.1 (ii), Module 5). Consequently $X_n^2 \xrightarrow{d} Q_1 \sim \chi_1^2$, as $n \rightarrow \infty$.
- (iii) Using Theorem 3.2 (ii) we get $X_n Y_n \xrightarrow{d} 3Z$, as $n \rightarrow \infty$. Let $V = 3Z$. Since $Z \sim N(0, 1)$ we have $V = 3Z \sim N(0, 9)$ (Theorem 4.2 (ii) Module 5) and, therefore, $X_n Y_n \xrightarrow{d} V \sim N(0, 9)$, as $n \rightarrow \infty$. Using theorem 3.2 (iii) and (iv) we get $2X_n \xrightarrow{d} 2Z$ and $3Y_n \xrightarrow{p} 9$, as $n \rightarrow \infty$. Now using Theorem 3.2 (ii) we also conclude that $2X_n + 3Y_n \xrightarrow{d} 2Z + 9$, as $n \rightarrow \infty$. Let $Q_2 = 2Z + 9$. Since $Z \sim N(0, 1)$, we have $Q_2 \sim N(9, 4)$ (Theorem 4.2 (ii), Module 5).
- (iv) From Example 1.4 we have $X_{n:n} \xrightarrow{p} \theta$, as $n \rightarrow \infty$, and $Y_n = n(\theta - X_{n:n}) \xrightarrow{d} Y \sim \text{Exp}(\theta)$, as $n \rightarrow \infty$. Since $h_4(x) = e^x$, $x \in \mathbb{R}$, $h_5(x) = x^2 + x + 1$, $x \in \mathbb{R}$, and $h_6(x) = e^{-\frac{x}{\theta}}$, $x \in \mathbb{R}$, are continuous functions on \mathbb{R} , using Theorem 3.1 (i) and (ii), we get $e^{X_{n:n}} \xrightarrow{p} e^\theta$, $X_{n:n}^2 + X_{n:n} + 1 \xrightarrow{p} \theta^2 + \theta + 1$ and $e^{-\frac{Y_n}{\theta}} \xrightarrow{d} e^{-\frac{Y}{\theta}}$, as $n \rightarrow \infty$. Let $U = e^{-\frac{Y}{\theta}}$. Since $Y \sim \text{Exp}(\theta)$, it is easy to verify that $U \sim U(0, 1)$. Consequently, $e^{-\frac{n(\theta - X_{n:n})}{\theta}} = e^{-\frac{Y_n}{\theta}} \xrightarrow{d} U \sim U(0, 1)$, as $n \rightarrow \infty$. ■